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# Geographically-targeted surveillance of livestock could help prioritize intervention against antimicrobial resistance in China 

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#### Abstract

The rise of antimicrobial resistance in animals is fueled by the widespread use of veterinary antimicrobials. China is the largest global consumer of veterinary antimicrobials - improving AMR surveillance strategies in this region could help prioritize intervention and preserve antimicrobial efficacy. Here, we mapped AMR rates in pigs, chickens and cattle in China using 446 surveys of event-based surveillance between 2000 and 2019 for foodborne bacteria, in combination with geospatial models to identify locations where conducting new surveys could have the highest benefits. Using maps of uncertainty, we show that eastern China currently has the highest AMR rates, and southwestern and northeastern China would benefit the most from additional surveillance efforts. Instead of distributing new surveys evenly across administrative divisions, using geographically targeted surveillance could reduce AMR prediction uncertainty by 2 -fold. In a context of competing disease control priorities, our findings


present a feasible option for optimizing surveillance efforts - and slowing the spread of AMR.

## Main

## Introduction

Antimicrobials are used in agriculture as disease treatments, prophylactically to prevent infections in healthy animals, and to increase productivity ${ }^{1}$. However, the routine use of antimicrobials as surrogates for good hygiene practices on farms ${ }^{2,3}$ is driving a rise in antimicrobial resistance (AMR), with increasingly serious consequences for animal health ${ }^{1,4}$, and potentially human health ${ }^{5,6}$.

Globally, $73 \%$ of antimicrobials are used in animals ${ }^{7}$, with China being the largest consumer of antimicrobials in absolute terms ( 41,967 tons in 2017) and the second largest consumer in relative terms with 200 milligram used per kilogram of animal $(\mathrm{mg} / \mathrm{kg})^{8}$ (Supplementary Figure 1A, 1B). In comparison, Denmark and the Netherlands use respectively $39 \mathrm{mg} / \mathrm{Kg}$, and $56 \mathrm{mg} / \mathrm{Kg}^{9}$ - while maintaining a productive livestock sector. Multiple factors may contribute to antimicrobial overuse in China. Meat production has grown by $560 \%$ since 1979 (FAOSTAT, http://www.fao.org/faostat/en/\#data/QL ), which could have made farmers reliant on antimicrobials to prevent infections. Veterinary antimicrobials are reportedly accessible without prescriptions ${ }^{10}$ and are sold at low prices in comparison to other countries ${ }^{11}$. As in many other low- and middle-income countries (LMICs) ${ }^{12,13}$, farmers predominantly obtain antimicrobials from local drug stores where vendors also provide medical advice
without veterinary training ${ }^{10,14}$. Additionally, enforcing the existing regulations ${ }^{10}$ on the compounds authorized in animals, or the recently announced ban on growth promoters ${ }^{15}$ remain a formidable challenge in a country where 360 million people are active in agriculture (World Bank, https://data.worldbank.org/indicator/SL.AGR.EMPL.ZS). In the last 5 years, China has reported multiple first emergence of resistance genes to last-resort antimicrobials such as colistin and tigecycline ${ }^{16,17}$ and a recent global analysis suggested that China may have become one of the largest hotspots of resistance among LMICs ${ }^{4}$, ranking $8^{\text {th }}$ in relative terms, and $1^{\text {st }}$ in absolute terms, for animal-associated burden of AMR amongst LMICs (Supplementary Figure 1D, 1E).

In high-income countries, epidemiological evidence collected by surveillance systems guides AMR responses and provides a baseline for evaluating policy targets. The US Food and Drug Administration collects meat samples from retail and slaughterhouses to monitor AMR levels (https://www.fda.gov/animal-veterinary/antimicrobial-resistance/national-antimicrobial-resistance-monitoring-system); the European Food Safety Authority (EFSA, https://www.efsa.europa.eu/en/topics/topic/antimicrobialresistance) serves a comparable role by amalgamating the surveillance efforts of its member states. To the best of our knowledge, the majority of LMICs - including China either lack systematic surveillance systems or do not publicly report data from animal AMR surveillance ${ }^{4}$. Despite these challenges, China could act as a leader for guiding the
international response to AMR - because its domestic policies may have far-reaching benefits for neighboring countries, and its numerous trading partners ${ }^{18}$.

Point-prevalence surveys (PPS) published independently by veterinarians constitute an alternative source for documenting AMR trends (Criscuolo et al. submitted) - and inferences can be made to map AMR using a large collection of PPS ${ }^{4}$. However, adapting this approach to the Chinese context requires building a critical mass of PPS, including surveys in Chinese to train geospatial models. Accurate maps of disease prevalence have been generated ${ }^{19-22}$, but few used the associated uncertainty maps to inform field sampling campaigns ${ }^{23,24}$. In particular, as prediction uncertainty grows with distance from existing surveys, an uncertainty map can help identify the location where conducting new surveys could be most valuable to improve the confidence level of a prevalence map. Repeating this process iteratively can guide long-term surveillance efforts.

Here, we used event-based surveillance data to map trends in AMR in animals and associated uncertainty levels. We identify regions where future surveillance efforts could be intensified to reduce uncertainty on the geographic distribution of AMR in China. In a
context of competing disease control priorities, our approach helps optimally target the limited resources dedicated to event-based surveillance of AMR.

## Results

## Data

We identified 446 point-prevalence surveys (PPS) reporting antimicrobial resistance (AMR) in food animals in China between 2000 and 2019 (Supplementary Text S1). This corresponds to one survey per 470,177 tons of food animals annually ( $28^{\text {th }}$ rank amongst low- and middle-income countries (LMICs); Supplementary Figure 1C). We collected data on four common indicator bacteria: Escherichia coli (184 PPS), nontyphoidal Salmonella spp. (131 PPS), Staphylococcus aureus (131 PPS), and Campylobacter spp. (33 PPS). The 446 PPS included 6,295 resistance rates. We defined a composite metric of AMR to summarize trends in resistance across multiple drugs, and bacteria. For each survey, we calculated the proportion of antimicrobial compounds with resistance higher than $50 \%$ (P50; Supplementary Figure 2).

## Temporal trends

In pigs, between 2000 and 2019, P50 increased significantly in E. coli $(+59 \%)$, Salmonella $(+148 \%)$, and $S$. aureus $(+85 \%)$ (Figure 1A, 1B, 1C). In contrast, in chicken, P50 was stable in E. coli, Salmonella, and S. aureus, with mean P50 of $60 \%, 42 \%$, and $37 \%$, respectively (Figure $1 \mathrm{D}, 1 \mathrm{E}, 1 \mathrm{~F})$. In cattle, P50 increased significantly in E. coli $(+167 \%$;

Figure 1G), and was stable in Salmonella and S. aureus, with mean P50 of $23 \%$ and $31 \%$, respectively (Figure 1H, 1I).

## Prevalence of resistance across antimicrobial classes

For each drug-bacteria-animal combination, we estimated the prevalence of resistance ( $\mathrm{R} \%$ ), and calculated the center of mass of the probability density distribution of the prevalence of resistance across PPS (Methods; Figure 2). Prevalence of resistance of tetracyclines, sulfonamides, and penicillins was high across all tested bacterial species between 2010 and $2019(\mathrm{R} \%>25 \%)$. In comparison, prevalence of resistance has remained at low levels in polymyxins and cephalosporins $(\mathrm{R} \%<10 \%$ for at least one bacterial species tested in one animal species). For all antimicrobial classes, prevalence of resistance in $E$. coli in chicken and pigs increased after 2010, except tetracyclines with already high prevalence of resistance ( $\mathrm{R} \%>90 \%$ ) in pigs before 2010. In Salmonella, increase in the prevalence of resistance after 2010 was observed in penicillins in chicken, as well as in sulfonamides, penicillins, and tetracyclines in pigs.

The prevalence of resistance in E. coli was higher than Salmonella for all antimicrobial classes (Figure 2A, 2B). Across drug classes, the prevalence of resistance in E. coli was $18 \%$ higher than the prevalence of resistance in Salmonella in chicken, and $16 \%$ higher than the prevalence of resistance in Salmonella in pigs. Prevalence of resistance for individual antimicrobial classes differed between chicken and pigs. For E. coli, cephalosporins and quinolones had respectively $20 \%$ and $27 \%$ higher prevalence of resistance in chicken compared with pigs, while prevalence of resistance in other
antimicrobial classes differed by $<6 \%$ between chicken and pigs (Figure 2A). For Salmonella, quinolones had $25 \%$ higher prevalence of resistance in chicken compared to pigs, while for other antimicrobial classes, the difference in the prevalence of resistance between chicken and pigs was smaller than $12 \%$. (Figure 2B). This comparison was largely influenced by the relative abundance of serotypes of Salmonella in different animal hosts (Supplementary Figure 3). However, an in-depth investigation on its influence on resistance trends was challenged by the fact that $70 \%$ of the surveys on Salmonella ( 93 out of 131 surveys) did not report the prevalence of resistance broken down by serotypes.

## Geographic Distribution of Resistance

We used a geospatial model (Supplementary Text S2) to map P50 at 10 Km resolution, and combined information from PPS with environmental and anthropogenic covariates (Supplementary Table 1). Hotspots of AMR - regions where more than $40 \%$ of drugs have resistance levels above $50 \%$ ( $\mathrm{P} 50>40 \%$ ) - were found in 1 ) eastern China in the areas of Heilongjiang, western Jilin, western Liaoning, southern Hebei, Shandong, eastern Jiangsu, southern Anhui, Fujian and Taiwan, 2) central China in the areas of northern Shaanxi, central Hunan and southeastern Sichuan, and 3) the northwestern Xinjiang Uyghur Autonomous Region (Figure 3). Low levels of AMR (P50 < 30\%) were found in Tibet Autonomous Region, northwestern Sichuan, and southern Guangxi (Figure 3). We measured the association between P50 and covariates, using the decrease in area under the receiver operator curve (AUC) by sequential permutation of each covariate (Supplementary Text S2). The most important covariates associated with P50 values were:
the travel times to cities ${ }^{25}(-16 \%$ AUC $)$, the minimum monthly temperature ${ }^{26}(-15 \%$ AUC), and cattle population density ${ }^{27}$ ( $-13 \%$ AUC; Supplementary Figure 4).

## Optimal location for future event-based surveillance efforts

We identified the locations of 50 hypothetical surveys to be conducted in China such that these would minimize uncertainty on the current map of AMR. The uncertainty was quantified using a map of "necessity for additional surveillance" $(N S)$ - the product of the kriging variance (a metric of interpolation uncertainty) and the population density (Methods). The 50 locations for the hypothetical surveys were identified such as to minimize the mean of $N S_{i}$ across all pixels $i$ in the $N S$ map.

We compared four approaches to distribute hypothetical future surveys (Methods): first a 'greedy' approach that tested all possible locations for new surveys but was associated with high computational cost. Second, an "overlap approach" based on mutual zones of exclusions for consecutive surveys to be conducted. This approach was a computational approximation to the greedy approach. Third, an 'administrative' approach where surveys were distributed equally across administrative divisions. These three approaches were compared with a "null model" consisting of randomly distributing 50 surveys across China. The greedy 'optimal' approach achieved the greatest reduction of the mean necessity for additional surveillance $(N S)$ (Figure 4B, dark red). The greedy approach reduced $N S$ by $56 \%$ more than the null model (Figure 4B, blue). However, the greedy approach was associated with a considerable computational burden (Figure $4 \mathrm{C}, 4.5 \times 10^{5} \mathrm{CPU}$ minutes). The overlap approach reduced the mean $N S$ by $44 \%$ more than the null model (Figure 4B,
blue), thus achieving near-optimal reduction of NS, but with a considerably lower computational burden than the greedy approach. The overlap approach also outcompeted the administrative approach (Figure 4B, green): it reduced the mean $N S$ by $104 \%$ more than if surveys had been distributed equally between administrative divisions.

The overlap approach predicted locations for a large number of new surveys in the southwest (21/50 surveys) and northeast (11/50 surveys) of China. The surveys were predominantly distributed in Yunnan Province (10 surveys), Tibet Autonomous Region (9 surveys), Xinjiang Uyghur Autonomous Region (7 surveys), and Heilongjiang Province (5 surveys) (Figure 4A). These locations were determined using animal population densities as the metric of exposure (Methods). Additionally, we calculated the locations by province for individual animal species, respectively (Supplementary Figure 5). The locations were mainly distributed in Yunnan and Heilongjiang (exposure by chicken or pigs), and in Tibet Autonomous Region (exposure by cattle). If human population was considered to determine exposure (Supplementary Figure 6), then the locations predicted by the overlap approach to conduct new surveys were mainly distributed in Heilongjiang Province (8 surveys), Xinjiang Uyghur Autonomous Region (8 surveys), Yunnan province (7 surveys), and Inner Mongolia Autonomous Region (5 surveys).

## Discussion

We identified geographical gaps in event-based surveillance of food animal AMR in China, using a map of AMR derived from 446 PPS, and its associated map of uncertainty,
identifying where surveillance scale-up would be the most valuable to reduce uncertainties in the current trends of AMR.

## Trends of AMR across animals and bacteria

Between 2000 and 2019, in pigs, P50 doubled in Escherichia coli, Salmonella, and Staphylococcus aureus. This increase in AMR occurred in a period of considerable intensification of pig production in China, and the number of pigs slaughtered in China increased by $45 \%{ }^{28}$. Traditional backyard systems were gradually replaced by large-scale intensive farms to support the growing domestic demand for pork ${ }^{29}$. However, as in other countries currently transitioning from extensive to intensive farming, improvements in biosecurity may have lagged behind improvements in productivity ${ }^{30}$. Future improvements in biosecurity may reduce farmers' dependency on antimicrobials for disease prevention, and have potentially indirect benefits for managing AMR in the longterm. Future biosecurity improvements can reduce the risk of diseases introduction through strict hygiene requirements for personnel who enter the farms, appropriate carcass management, and reducing the spread of diseases inside the premises through establishing pig compartments, and regular cleaning and disinfection ${ }^{31}$.

Between 2000 and 2019, in chicken, P50 remained stable in E. coli, Salmonella, and $S$. aureus, albeit at high levels. In 2000, P50 in E. coli, Salmonella, and S. aureus in chicken were already at $58 \%, 48 \%$, and $56 \%$, respectively, double the levels of resistance in pigs $(35 \%, 15 \%$, and $30 \%)$. This suggests that the intensification process (and the routine use of antimicrobials for production) occurred earlier and faster in the poultry sector than for
pigs ${ }^{32}$. Excessive use of quinolones (e.g. norfloxacin and ofloxacin) and cephalosporins (e.g. ceftriaxone) in chicken ${ }^{10}$ may have caused much higher resistance rates of these two antimicrobial classes in chicken, compared with pigs (Figure 2). Our analysis suggests that the antimicrobials that maintained low prevalence of resistance in chicken are expensive and are seldom available on the Chinese market (Supplementary Text S3), impeding overuse and also preventing the further AMR increases.

Prevalence of resistance for E. coli were higher than for Salmonella in pigs and chicken (Figure 2), possibly influenced by commensal E. coli being associated with lower resistance levels than pathogenic E.coli ${ }^{33}$. However, due to the non-systematic nature of the PPS sampling schemes (event-based surveillance), disentangling how resistance rates differ between bacteria exhibiting commensal or pathogenic behavior remains challenging. We attempted to mitigate this potential bias by focusing our analysis exclusively on bacteria isolated from healthy animals.

Resistance levels (P50) in cattle were lower than in chicken and pigs (Figure 1). However, P50 in E. coli grew by $81 \%$ between 2000 and 2019, while globally the P50 in cattle was stable over the same period ${ }^{4}$. This may be associated with the increasing demand for cattle product in China - cow milk production increased by $261 \%$ from 2000 to $2019{ }^{28}$. Despite this rapid expansion, the current per capita consumption of dairy products in China is still only one fifth of the dairy consumption in the US and the EU ${ }^{28}$ - leaving room for further expansion. Thus, a window of opportunity may exist at the
current stage to slow the rise of AMR in cattle, while resistance rates are still low ( $22 \%$ in E. coli) - and immediate action could help secure a sustainable dairy intensification.

## Improved maps of AMR in China

Currently, AMR levels in animals are the highest in the east (43\%), moderately high in the northwest (40\%), and lowest in the southwest (34\%; Figure 3). These geographical trends are in agreement (Pearson correlation coefficient 0.48 ) with the previous attempts to map AMR in China ${ }^{4}$. However, the present map is considerably more robust because it is exclusively based on surveys conducted in China (446 surveys, including 318 publications in Chinese). In comparison, previous maps were produced with just 101 surveys from China supplemented by surveys from other LMICs ${ }^{4}$. The revised maps of AMR help identify hotspots of AMR (Figure 3) where intervention could be targeted immediately as part of domestic policies ${ }^{34}$. Travel time to cities was the factor with the highest influence on resistance levels ${ }^{25}$. The clustering of intensive farms in major consumption centers during industrialization ${ }^{35}$, and the ease of access to drug stores in peri-urban areas ${ }^{10,36}$ may drive AMR level upwards ${ }^{37}$. High AMR levels were also associated with high minimal monthly temperature ${ }^{26}$ - high temperatures may lead to increased stress and conflicts among animals, with risk of animal injuries requiring antimicrobial treatment ${ }^{38}$.

## Key locations for conducting event-based surveillance

Amongst LMICs, China ranks $28^{\text {th }}$ for the number of surveys in event-based surveillance per kilogram of food animals (population corrected units of food animals; PCU), and $36^{\text {th }}$ for the number of surveys per PCU relative to average resistance level (P50) per country
(Supplementary Figure 1). We identified locations where additional surveys on AMR in animals could be conducted in the future to minimize uncertainty associated with the geographical trends in AMR - representing a gain in information given the resources spent on event-based surveillance. Current patterns resulting from event-based surveillance are "sub-optimal" as surveys are clustered around veterinary institutes, mainly in the east (Supplementary Figure 7), where sampling to investigate AMR in their vicinity is easier (Supplementary Text S4) - and may have contributed to geographical information gaps on AMR trends in the southwest and northeast. Cross-provincial efforts between institutes are needed to coordinate future event-based surveillance efforts into these regions, which may be far from existing institutes, but where the gain in information by additional surveys would be the highest.

Our approach for assigning future surveys works by minimizing an index of "need for additional surveillance (NS)" based on a map of AMR (China for this example). However, testing exhaustively all possible location for future surveys ("greedy approach") incurs considerable computational cost. We developed an 'overlap approach' which is a rapidly implementable approximation of a greedy approach. The overlap approach achieved $93 \%$ of the reduction of the uncertainty in AMR trends achieved by the greedy approach, albeit using just $15 \%$ of total computation time required by the greedy approach (Methods). This not only makes the approach faster but also applicable with limited computational resources, and was developed in the context of event-based surveillance, which was abundant in China with 446 PPS and served as a proof of concept. However, the approach could also be used with systematic surveillance data or
in other countries with event-based surveillance. In addition, the 'overlap approach' is flexible with respect to exposure. In this analysis we used animal densities as metric of exposure, but this variable could easily be substituted by other criteria that are relevant for epidemiological or environmental assessments.

## Limitations and Future Directions

Although steps are taken (Supplementary Text S1) to ensure comparability between surveys, there remain potential sources of bias in variations in the accuracy of susceptibility testing. These include potential difference in laboratory equipment, and compliance to analysis protocols across regions in China. The World Health Organization assesses the quality of antimicrobial susceptibility testing across countries ${ }^{39}$, but to the best of our knowledge, such within-country assessment is not currently available to account for laboratory practices that could lead to variations in the accuracy of susceptibility testing. These 'hidden' variations between surveys may influence the accuracy of the spatial distribution of P50. Inherent to event-based surveillance, a subjective summary metric "P50" was used in the absence of publicly available systematic surveillance data. P50 could be affected by the different antimicrobials subject to susceptibility testing in each survey. The potential bias was reduced by using the drugbacteria combinations recommended by the WHO Advisory Group on Integrated Surveillance of Antimicrobial Resistance ${ }^{40}$ to calculate P50. Insufficient and irregular geographic coverage of data points may affect the accuracy of the estimations of model parameters. The risk of local overfitting is attenuated by using spatial cross-validation in the models. Finally, future mapping efforts could integrate surveys on AMR in
aquaculture, because aquatic animals are important food animals in China, with at least 20 antimicrobials involved in their production ${ }^{41}$. Complementary to phenotypic resistance, AMR surveillance could be expanded to include genomics data, through metagenomic analysis of wastewater ${ }^{42}$ from farms, although issues about harmonization remain an active field of analysis (Pires et al. submitted).

The health challenges that China currently faces are multifaceted and burdensome, both in humans (e.g. COVID-19 ${ }^{43}$ ), and in food animals (e.g. African Swine Fever ${ }^{44}$ ). With limited resources to allocate between competing priorities for disease surveillance, our approach identifies locations where conducting new surveys of AMR in animals could have the highest benefits, particularly in southwestern and northeastern China. Timely policy intervention could curb AMR in China, as illustrated by the significant reduction in colistin resistance after the colistin withdrawal policy ${ }^{45}$. Our analysis helps to optimally deploy the limited resources dedicated to event-based surveillance of AMR improving chances for successful intervention for curbing AMR development and providing data to inform policy.

## Methods

## Data

We reviewed point-prevalence surveys (PPS) reporting rates of antimicrobial resistance (AMR) in healthy animals, and animal food products in China between 2000 and 2019 (Supplementary Text S 1 ). We focused on three common food animal species, including chicken, pigs, and cattle. Here, dairy cattle and meat cattle were pooled in this study, in
consistency with the categorization adopted in the maps of livestock created by the Food and Agriculture Organization ${ }^{27}$. The review focused on four common foodborne bacteria: Escherichia coli, nontyphoidal Salmonella, Staphylococcus aureus and Campylobacter. We recorded resistance rates reported in PPS, defined as the percentage of isolates tested resistant to an antimicrobial compound. In addition, we extracted the anatomical therapeutic chemical classification codes of the drugs tested, the year of publication, the guidelines used for susceptibility testing, the latitude and longitude of sampling sites, the number of samples collected, and the host animals. We recorded sample types for each survey, including live animals, slaughtered animals, animal products, and fecal samples. Each sample was taken from one animal or animal product. These sample types were pooled in the current analysis. 10,747 rates of AMR were extracted from 446 surveys (Supplementary Figure 8), including 318 surveys from China's National Knowledge Infrastructure (CNKI), the leading Chinese-language academic search engine. All data extracted in the review are available at https://resistancebank.org.

Two steps were taken to ensure comparability of the resistance rates extracted from the surveys. First, the panel of drug-bacteria combinations extracted from each survey was that recommended for susceptibility testing by the WHO Advisory Group on Integrated Surveillance of Antimicrobial Resistance (AGISAR) ${ }^{40}$. This resulted in the extraction of 6,295 resistance rates for 76 drug-bacteria combinations. Second, resistance rates were harmonized using a methodology ${ }^{4}$ accounting for potential variations in the clinical breakpoints used for antimicrobial susceptibility testing (Supplementary Text S1). There are two major families of methods used for susceptibility testing in this dataset-diffusion
methods (e.g. disc diffusion) and dilution methods (e.g. broth dilution). Previous works have shown good agreement between the two approaches in measuring resistance in foodborne bacteria ${ }^{4,46}$. For each family of methods, variations of breakpoints may result from differences between laboratory guidelines systems (European Committee on Antimicrobial Susceptibility Testing; EUCAST vs Clinical and Laboratory Standards Institute; CLSI), or from variations over time of clinical breakpoints within a laboratory guidelines system (CLSI or EUCAST). Here, we accounted for both situations using distributions of minimum inhibitory concentrations, and inhibition zones obtained from eucast.org (Supplementary Text S1).

## Trends in Antimicrobial Resistance

We defined a composite metric of AMR to summarize trends in resistance across multiple drugs and bacterial species. For each survey, we calculated the proportion of antimicrobial compounds with resistance higher than $50 \%$ (P50). For each animalbacteria combination, we assessed the significance of the temporal trends of P50 between 2000 to 2019 using a logistic regression model, weighted by the $\log 10$ transformed number of samples in each survey.

For each bacteria-drug (antimicrobial class) combination, we estimated prevalence of resistance by calculating a curve of the distribution of resistance rates across all surveys (Figure 2). The analysis was conducted for surveys published between 2000 and 2009, and between 2010 and 2019, respectively. The distribution was estimated at 100 equally spaced intervals from resistance rates of $0 \%$ to $100 \%$, using kernel density estimation.

We used the center of mass of the density distribution to estimate prevalence of resistance. The calculation was conducted for six animal-bacteria combinations. This included E. coli in chicken, pigs and cattle, as well as Salmonella in chicken and pigs, and $S$. aureus in cattle. The remaining animal-bacteria combinations were excluded due to limited sample size, only represented in 32 out of 446 PPS. The analysis was restricted to antimicrobial classes represented by at least 10 resistance rates. In addition, we estimated the association between resistance rates and the ease of obtaining antimicrobials from the market, using data from online stores (Supplementary Text S3).

## Geospatial modelling

We interpolated P50 values from the survey locations to create a map of P50 at a resolution of $10 \times 10 \mathrm{Km}$ across China. The approach followed a two-step procedure ${ }^{47}$ : In step 1, three 'child models' were trained using four-fold spatial cross validation to quantify the relation between P50 and environmental and anthropogenic covariates (Supplementary Text S2; Supplementary Table 1). In step 2, the predictions of the child models were stacked using universal kriging (Supplementary Text S2). This approach combined the ability of the child models to capture interactions and non-linear relationships between P50 and environmental and anthropogenic covariates, as well as the ability to account for spatial-autocorrelation in the distribution of P50.

The outputs of the two-step procedure were: a map of P50 (Figure 3), and a map of uncertainty on the P50 predictions (Supplementary Figure 9, Supplementary Text S2). The overall accuracy of the geospatial model was evaluated using the area under the
receiver operator curve (AUC). The contribution of each covariate was evaluated by permuting sequentially all covariates, and calculating the reduction in AUC compared with a full model including all covariates (Supplementary Figure 4). The administrative boundaries used in all maps were obtained from the Global Administrative Areas database (http://www.gadm.org).

## Identifying (optimal) locations for future surveys on AMR

We identified the locations of 50 hypothetical new surveys - the rounded average number of surveys conducted per year (54 surveys/year) between 2014 and 2019 in China. The location of each new survey was determined recursively such that it minimized the overall uncertainty levels on the geographical trends in AMR across the country. This process took into account the locations of existing surveys, as well as the location of each additional hypothetical survey. The objective of this approach was to maximize gain in information about AMR given the resource invested in conducting surveys.

The map of uncertainty consisted of the variance in the child model predictions $\operatorname{Var}\left(P_{B R T}, P_{L A S S O-G L M}, P_{F F N N}\right)($ step 1) across 10 Monte Carlo simulations, and the kriging variance $\operatorname{Var}_{K}$ (step 2):

$$
\operatorname{Var} r_{\text {total }}=\operatorname{Var}\left(P_{B R T}, P_{L A S S O-G L M}, P_{F F N N}\right)+V a r_{K}
$$

In this study, the location of hypothetical surveys was solely based on $V a r_{K}$, instead of the sum of both terms. This approach was preferred, because including both terms would
have required to hypothesize P50 values associated with the surveys to be conducted in the future, adding an additional source of uncertainty that cannot be quantified. In any case, the uncertainty attributable to $V a r_{K}$ was 4.1 times of $\operatorname{Var}\left(P_{B R T}, P_{L A S S O-G L M}, P_{F F N N}\right)$ (Supplementary Text S2).

The allocation of new surveys was based on a map of "necessity for additional surveillance" $(N S)$, defined as:

$$
N S=V a r_{K} \cdot W
$$

Where $V a r_{K}$ reflected the uncertainty of the spatial interpolation, and $W$ is $\log 10$ transformed population density of humans ${ }^{48}$, animals ${ }^{27}$ in total, as well as in chicken, pigs, and cattle, separately, which reflected exposure (Supplementary Figure 10). Here, animal population density was calculated as the sum of population corrected units (PCU) of pigs, chicken and cattle, using methods described in Van Boeckel et al. $2017^{7}$. We adjusted the values of $W$, such that its density distribution equals that of $V a r_{K}$. Concretely, for each pixel $i$, we calculated the quantile of $W i$ on the map of $W$, and replaced the value by the corresponding value of $V a r_{K}$ at the same quantile. $V a r_{K}$ and $W$ were both standardized to range [0,1], thus given each term equal weight in the need for surveillance.

Four approaches were used to distribute 50 surveys across China based on the map of $N S$. The reduction in uncertainty on AMR level associated with each of the four spatial
configurations of the hypothetical surveys was evaluated, by calculating the reduction in the mean values of $N S$ across 7,857 possible pixels on the map of China.

First, we used a "greedy" approach where all possible locations for additional surveys were tested. Concretely, the first hypothetical survey was placed at each of the 7,857 possible pixel locations, and a revised map of $N S_{(+1 \text { survey })}$ was calculated for each of the placements. The survey was eventually placed in the pixel that led to the largest reduction in $N S_{(+1 \text { survey) }}$. The map of $N S$ was then revised to account for the reduction in uncertainty in the neighborhood of the new survey. The process was repeated recursively for the next hypothetical surveys ( $\left.2^{\text {nd }}-50^{\text {th }}\right)$. This approach, by definition, yields the optimal set of locations to reduce uncertainty, but it also bears a considerable computational burden, because every possible location is tested $\left(N_{\text {pixels }}=7,857\right)$ by the geospatial model for each hypothetical survey.

The second approach developed was a computational approximation to the greedy approach, hereafter referred to as the 'overlap approach'. This approach exploits a key feature of the kriging procedure: the decrease of the kriging variance $\left(V a r_{K}\right)$ with increasing proximity to existing survey locations. Each additional survey reduces the variance of the geospatial model at its own location, but also in its surrounding area (Supplementary Figure 11). The 'overlap approach' selects an optimal set of locations that reflect a compromise between high local $N S$ and distance to other surveys. It iteratively selects new locations based on the highest local $N S$ penalized by the degree of
overlap between the hypothetical new surveys and existing surveys (Supplementary Figure 12). The first survey was placed at the location $X_{p}, Y_{p}$ with the highest local $N S$ (Supplementary Figure 12, Part 1). Then, the value of $N S$ at each pixel location $X_{i}, Y_{i}$ was recalculated as (Supplementary Figure 12, Part 2):

$$
N S_{\left(+1 \text { survey) } X_{i}, Y_{i}\right.}=N S_{X_{i}, Y_{i}} \times\left(1-\frac{\text { overlap area }}{\text { neighborhood area }}\right)
$$

Where the neighborhood area was the circular area of decreased kriging variance around a new survey, and its radius was the distance until which $N S$ decreased due to this new survey; "Overlap area" is the shared area of the neighborhoods of location $X_{p}, Y_{p}$ and of location $X_{i}, Y_{i}$. The radius of the neighborhood was determined using a sensitivity analysis, optimized by approximate Bayesian computation (sequential Monte Carlo) ${ }^{49}$ (ABC-SMC; Supplementary Text S5). The optimal neighborhood radius was chosen such as it minimizes reduction in $N S$ across all pixels. The procedure (Supplementary Figure 12, Part 1 and Part 2) was repeated recursively for the hypothetical surveys $\left(2^{\text {nd }}-50^{\text {th }}\right)$.

The third approach tested consisted of distributing surveys equally between provinces, to reflect a common approach to disease surveillance based on equal allocation of resources between administrative entities. Here, 22 provinces with the highest human population were assigned 2 surveys, and the remaining 6 provinces were assigned one survey per province. The exact location of each survey was randomly selected inside a province.

Finally, all approaches were compared with the fourth approach (the random approach) as a 'null-model', in which the 50 hypothetical surveys were located randomly across the country without any geographic weighting criteria. The reduction in $N S$ associated with the third and fourth approach, which was compared to the greedy approach and overlap approach, was the average over 50 simulations.

## Data Availability

Resistance rates and related information extracted from all point-prevalence surveys can be accessed in the Supplementary Data file, and at https://resistancebank.org.

## Code Availability

The codes used to generate the results are available at Zenodo (https://doi.org/10.5281/zenodo.4940089).

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## Author Contributions

CZ, TVB and JP developed the analytical framework. CZ, YW and KT conducted the literature review. CZ conducted the analysis and wrote the first version of the manuscript.

NC incorporated all data into resistancebank.org. All authors contributed to the final version of the manuscript.

## Competing interests:

No competing interests declared.

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Figure 1. Antimicrobial resistance between 2000 and 2019. P50 is the proportion of antimicrobials with resistance higher than $50 \%$ (P50). Mean refers to the mean P50 value of all surveys. $\mathrm{C}_{1}$ is the coefficient associated with the temporal trend in a logistic regression model weighted by $\log 10$ transformed sample size in each survey. Shaded area indicates $95 \%$ confidence intervals. ${ }^{* * *} \mathrm{p}<0.001$; ${ }^{* *} \mathrm{p}<0.01$; ${ }^{*} \mathrm{p}<0.05 . \mathrm{N}_{\text {surveys }}$ is the number of surveys, and $\mathrm{N}_{\text {rates }}$ is the total number of resistance rates reported in the surveys. Surveys conducted at multiple locations in the same publication are considered multiple surveys.


Figure 2. Prevalence of resistance per antimicrobial class. In each panel, $x$-axis represents resistance rates, and y-axis represents the probability density. (A) Escherichia coli in chicken, pigs and cattle; (B) Salmonella in chicken and pigs; (C) Staphylococcus aureus in cattle. The x-axis represents resistance rates, and the area under the curve between two resistance rates represents the probability for resistance rates to fall within the interval. N : number of resistance rates used to calculate the density distribution. Dashed lines represent the center of mass of each distribution.


Figure 3. Geographic distribution of antimicrobial resistance. Color represents the proportion of antimicrobials with resistance higher than $50 \%$ (P50).


Figure 4. Predicted Locations for Future Surveys. (A) Predicted optimal locations for future surveys using the "overlap approach". The background color represents the "necessity for additional surveillance" $(N S)$ : the product of the kriging variance and animal population density (standardized from 0 to 1 ). (B) Reduction in the mean $N S$ with 50 hypothetical additional surveys. The 50 additional survey locations were identified using the greedy approach (dark red), the overlap approach (red), the administrative approach (green), and the random approach (blue). (C) Total CPU (central processing unit) time for computing the four approaches ( $\log 10$ scaled).

